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ON THE REPRESENTATION OF THE
DENSITY FUNCTIONS GAMMA TYPE

To Professor Lech Włodarski on His 80th birthday

In this paper we present the random variable X of gamma type as infinite (or finite) product of independent random variables X_k , $k \in N$, in the following sense:

$$X \stackrel{st}{=} \prod_{k=1}^{\infty} X_k.$$

Let the random variable X have three-parameters gamma distribution with the density given by the formula

(1)

$$f(x|\Theta, \lambda, p) = \frac{\lambda}{\Theta^{p/\lambda} \Gamma^{-1}\left(\frac{p}{\lambda}\right)} x^{p-1} \exp\left(-\frac{x^\lambda}{\Theta}\right), \quad x \geq 0, \quad p, \lambda, \Theta > 0,$$

where $\Theta^{1/\lambda}$ is the scale parameter and p, λ are the shape-parameter.

In the present note it will be used the Mellin transform (2) of the function (1)

$$(2) \quad \begin{aligned} M(s) &= M[f]x|\Theta, \lambda, p, s] = EX^s \\ &= \Theta^{s/\lambda} \Gamma\left(\frac{p+s}{\lambda}\right) \Gamma^{-1}\left(\frac{p}{\lambda}\right), \quad \text{Re } s > -p, \end{aligned}$$

where s denote a complex variable: [1], [2]: recall that Sneddon [3], Dytkin and Prudnikow [4] have defined the Mellin transform $M(s)$ as EX^{s-1} .

By applying the Knar formula [5.8, p. 324]

$$\Gamma(x+1) = 2^{2x} \prod_{k=1}^{\infty} \frac{\gamma\left(\frac{1}{2} + \frac{x}{2^k}\right)}{\Gamma\left(\frac{1}{2}\right)}, \quad \operatorname{Re} x > -1,$$

into (2) we obtain

$$M(s) = (4\Theta)^{s/\lambda} \prod_{k=1}^{\infty} \frac{\gamma\left(\frac{1}{2} + \frac{p-\lambda+s}{2^k\lambda}\right)}{\gamma\left(\frac{1}{2} + \frac{p-\lambda}{2^k\lambda}\right)}, \quad \operatorname{Re} s > -p.$$

The last formula can be written as

$$(3) \quad M(s) = \prod_{k=1}^{\infty} \frac{(4\Theta)^{s/\lambda 2^k} \Gamma\left(\frac{1}{2} + \frac{p-\lambda+s}{2^k\lambda}\right)}{\Gamma\left(\frac{1}{2} + \frac{p-\lambda}{2^k\lambda}\right)}, \quad \operatorname{Re} s > -p.$$

Let us denote k -factor occurring in the right hand side of (3) by $g_k(s)$. Next, we apply the inverse Mellin transform for every factor $g_k(s)$; the result of this transform is denoted by $f_k(x|\Theta, \lambda, p)$:

$$(4) \quad f_k(x|\Theta, \lambda, p) = \frac{1}{2\pi i} \int_{c-i\infty}^{c+i\infty} x^{-s-1} g_k(s) ds, \quad \operatorname{Re} s > -p.$$

Putting $c = 0$ in view $\operatorname{Re} s > -p$, we get

$$(5) \quad f_k(x|\Theta, \lambda, p) = \frac{2^k \lambda}{\Gamma\left(\frac{1}{2} + \frac{p-\lambda}{2^k\lambda}\right) (4\Theta)^{\frac{1}{2} + \frac{p-\lambda}{2^k\lambda}}} x^{(2^{k-1}-1)\lambda+p-1} \exp\left(-\frac{x^{2^\lambda}}{2\Theta}\right), \quad x \geq 0.$$

It is easy to prove that the condition $\int_0^\infty f_k(x|\Theta, \lambda, p) dx = 1$ is satisfied for any $k \in \mathbb{N}$. In view of the inequality $f_k(x|\Theta, \lambda, p) \geq 0$

this implies that the formula (5) defines the density of probability of the random variable X_k in the interval $[0, \infty)$. In consequence every factor $g_k(s)$, $k \in \mathbb{N}$ of the infinity product (3) is the Mellin transform of the random variable X_k with the density given by (5).

Now we shall show that $\prod_1^\infty g_k(s)$ is the Mellin transform of $\prod_1^\infty f_k(x|\Theta, \lambda, p)$. To do this we find an estimate of an absolute value of any factor of (3). In fact setting $\operatorname{Re} s = c = 0$ according to $\operatorname{Re} s = c = 0$ we get

$$\left| \frac{(4\Theta)^{s/2^k \lambda} \Gamma\left(\frac{1}{2} + \frac{p-\lambda+s}{2^k \lambda}\right)}{\Gamma\left(\frac{1}{2} + \frac{p-\lambda}{2^k \lambda}\right)} \right| \leq \frac{(4\Theta)^{\operatorname{Re} s/2^k \lambda} \Gamma\left(\frac{1}{2} + \frac{p-\lambda+\operatorname{Re} s}{2^k \lambda}\right)}{\Gamma\left(\frac{1}{2} + \frac{p-\lambda}{2^k \lambda}\right)} \leq 1.$$

Let us denote $M_n((s)) = \prod_{k=1}^n g_k(s)$. Besides, for any $n \in \mathbb{N}$ $|M_n(s)| \leq |M_1(s)|$ and $M_1(s)$ is an absolute integrable function. From the above and in view of the Lebesgue dominated convergence theorem [6. Th. 1.34] and (3) we obtain successively

$$\lim_{n \rightarrow \infty} \int_{-i\infty}^{i\infty} x^{-s-1} \left(M(s) - \lim_{n \rightarrow \infty} M_n(s) \right) ds = 0.$$

Thus, we have proved the following

Theorem. *The density of the three-parameter gamma distribution of the form (1) is equal to the density of the infinite product $\prod_1^\infty X_k$, where the density $f_k(x|\Theta, \lambda, p)$ of X_k are independent and defined by the formula (5), the scale parameter being $(4\Theta)^{1/2^k \lambda}$ respectively, what can be formally denoted as*

$$X \stackrel{\text{st}}{=} \prod_{k=1}^\infty X_k.$$

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O PEWNEJ REPREZENTACJI FUNKCJI GĘSTOŚCI TYPU GAMMA

Przed ca 30 laty ukazało się wiele prac dotyczących wyznaczania gęstości skończonego iloczynu zmiennych losowych niezależnych o rozkładach Beta, Gamma, Normalnym, Bessela, w tym prace pracowników naukowych Politechniki Łódzkiej: Śródki, Podolskiego i Krysickiego.

Prezentowana praca stawia sobie jako cel: cel przeciwny, którym jest przedstawienie zmiennej losowej X typu Gamma jako nieskończonego (bądź skończonego) iloczynu zmiennych losowych niezależnych X_k , $k \in \mathbb{N}$, s sensie stochastycznym tzn. w sensie równości gęstości $X \stackrel{\text{stoch}}{=} \prod_1^\infty X_k$.

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